

Decision Support for Service Transition Management

Enforce Change Scheduling by Performing Change Risk and Business Impact Analysis

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Abstract In IT Service Delivery, alignment of service infrastructures to continuously changing business requirements is a primary cost driver, all the more as most severe service disruptions can be attributed to poor change impact and risk assessment. An IT service, defined as a means to provide value to a consumer, may be realized by a network of shared application and other resources that are invoked in the context of business processes. In the spirit of Service-Oriented Architecture (SOA) we consider each application or resource as a service. Changing services or service definitions in such an environment includes exceptionally high risk and complexity, as various business processes might depend on a service. In this paper we propose a model for analyzing the business impact of operational risks resulting from change related service downtimes of uncertain duration. The proposed solution takes into account the network of dependencies between services where services may or may not be realized through business processes. Based on the analytical model, we derive decision models in terms of deterministic and probabilistic mathematical programming formulations to schedule single or multiple correlated changes efficiently. Preliminary experiments are described to illustrate the efficiency of the proposed models. Using these decisions models, organizations can schedule service changes with the lowest expected impact on the business.

Change Management, Service Transition Management, Change Scheduling, Service-oriented Architectures, Business Impact Analysis, Business-Driven IT Management

I. INTRODUCTION

In recent years, IT service management (ITSM) has received much attention as enterprises understand that operating their IT infrastructure is a large part of their overall operating costs. Today's businesses operate in dynamic environments with the need to continuously adapt to changing customer expectations, market trends, technical enhancements or changes to legislation. These changes entail changes to IT services and business processes to drive alignment of IT with business requirements. According to current surveys uncontrolled changes including flawed risk and impact analysis cause more than 80% of business-critical service disruptions [1].

Publicly available best-practices ITSM frameworks such as the IT infrastructure Library (ITIL) define reference change management processes including several activities like change initiation, where a Request for Change (RFC) describing the

required change is submitted, change filtering, priority allocation, categorization, planning, testing, fulfillment and review. Major changes must be analyzed and approved, from a technical as well as from a business point of view before they get scheduled [2]. We focus on the impact of changes on the business and on how to schedule changes with minimum associated risks and costs.

As modern IT service infrastructures are continuously transformed towards virtualized resource pools and Service-Oriented Architectures, applications and infrastructure resources can be viewed as services shared in a larger value network and invoked in the context of various business processes. Services can be described using standards such as the Web Services Description Language (WSDL), describing for example the service interfaces of services offered as Web services, and invoked via a suitable Internet protocol. We distinguish between different types of services. An atomic service in our definition is a service with a well-defined transaction boundary that provides a simple single operation (e.g. *generate IP* or *assignServerName*). A business process executes by invoking atomic services, other services that may be composed of atomic services (e.g. short running automated workflows) or other business processes. Each service is executed on an IT resource. In principle, we can consider an IT resource as a service as well. We exclude the details of the implications of this approach for a later time.

Considering the number of business processes in an enterprise and the complexity of the dependency network of processes to invoked services, changes in this kind of environment may pose significant risks due to the multitude of interdependencies and uncertainties to manage, and the impact of failures is likely to be business-critical as many business processes might depend on this service. Therefore, efficient and reliable change management aiming at continuous service delivery by automatically considering the dependency chains is essential.

Consider the following example, illustrated in Fig. 1, which we found in many companies: a business process application runs several customer relationship management (CRM) processes, from lead generation to sales order generation (CRM covers concepts to manage relationships with customers, including the capture, storage and analysis of customer, vendor, partner, and internal process information). The application itself is hosted on one or more physical resources and has

dependencies to other applications (or services) and infrastructural components. Estimating the impact of an application failure is – without detailed knowledge of the dependency chains - a fairly manageable problem. The left scenario shows a pictorial of a *Tivoli CCMDB discovery* [3], where application A is connected to application B. Downtime of Application B means an impact on Application A. This view however is not sufficient as an organization managing the business process application A will alert business users that the CRM application will be unavailable, which could for example lead to unfulfilled sales orders. The right pictorial illustrates the more realistic scenario, where A is hosting two processes, e.g. lead generation and sales order generation. The actual downtime of application B may only lead to unavailability of lead generation but not sales order generation (which in the CRM context is a much lower risk). Furthermore, based on the fact that the affected lead generation may be a long-running business process, one can imagine that only a subset of all running instances will be affected depending on the state of each instance. The longer the duration of downtime for a given service or application or network resource that is used by a business process is, the more likely it is to experience business value attrition due to SLA violations and associated penalties. We believe that change scheduling should take this level of detail into account to minimize the risk of downtime for business value generating services.

How many instances of a particular process are affected highly depends on the business process demand while fulfilling the change. Process demand, however, is generally not known a-priori but has to be approximated by means of forecasting techniques.

The focus of this paper is to determine and minimize change related risk in Service-Oriented Business environments as illustrated above by introducing decisions models allowing organizations for scheduling service changes with the lowest expected financial loss, or cost. We propose models for analyzing the business impact of change related service downtimes of uncertain length, as the impact on dependent, active business processes is analyzed and transferred into financial losses. The proposed solution automatically considers the dependency chain from a business process down to affected resources, applications or other services realized by business processes. Based on these analytical models, we derive decision models in terms of deterministic and probabilistic mathematical programming formulations allowing for scheduling single or multiple correlated changes efficiently. First Experiments and sensitivity analyses are described to illustrate the efficiency of the proposed models.

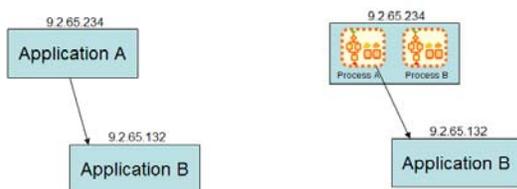


Figure 1. Business Process Application Dependencies

The outline of this paper is structured as follows: in Section 2 we review related work in this field. In Sections 3 we discuss in detail techniques on how to estimate and quantify operational risks of service transitions. Subsequently, in Section 4 we introduce a basic deterministic decision model and probabilistic extensions to determine efficient change schedules in different business scenarios. We further provide model extensions to take into account change correlations and other sources of risks like change deadline or change window violation risks. In Section 5, we describe the setup of our preliminary experiments conducted to make first efficiency statements of the models and present our experimental outcomes. Finally, in Section 6, we summarize and conclude.

II. RELATED WORK

In this section we review previous work and guidelines in IT change risk analysis and management related to the work described in this paper.

The definition of risk itself varies broadly according to the specific domain one looks at. The most general definition of risk is ‘uncertainty of outcome’ [4]. In our case, the outcome is change related cost materialized as financial loss. To analyze risk impact, i.e., resulting costs for the business, we draw on a two-stage approach; first scanning for possible outcomes and quantifying this outcomes in terms of monetary consequences, and second, weighting these outcomes by their probabilities. This approach is generally known as probabilistic risk analysis as introduced in [5].

In IT change management, most approaches found allow for risk analyses that are not directly transferable to business or financial impact or provide general guidance to change management considering associated risk.

Publicly available best-practices ITSM frameworks and standards such as the IT Infrastructure Library (ITIL) [6] or Control Objectives for Information, and related Technology (COBIT) [7] provide guidance on how to perform service management tasks and are validated across a diverse set of environments and situations. As of the importance of managing service changes or transitions efficiently, particular with respect to associated risks, this topic has recently become a mayor focus herein. For example, the Office of Governmental Commerce (OGC) dedicated in the newly published ITIL version 3.0 (May 2007) an own book on how to manage service transitions efficiently, with special regards to associated risks [4]. However, ITIL and related best-practices frameworks provide high-level guidance for performing a service management task like managing a change, but do not provide guidance on how to do the actual change management implementation, e.g., on how to determine and quantify change related risks and costs in a particular business environment.

Some commercial tools and dashboard applications are available that claim to assist in managing changes, although not enough details are available that can be used to evaluate and compare the involved methods [8, 9, 10, 11].

Some papers have presented approaches to qualitatively evaluate risk, for example [12], but do not provide quantitative risk analysis with regard to business impact.

Keller and Hellerstein present the CHange Management with Planning and Scheduling (CHAMPS) system to automate steps in the execution of changes. The authors propose decision models to solve different scheduling problems like maximizing the number of changes, minimizing overall downtime, or minimizing the costs associated with change related downtime. The authors assume knowledge of the cost functions for performing a change job at time t , while we focus on how to derive cost functions from change related downtime risks to the business processes [13].

Rebouças, Sauv , Moura, Bartolini and Trastour address the problem of scheduling changes in a way to minimize the financial loss imposed by SLA violations when the implementation of changes exceed change deadlines. The authors explicitly consider uncertainty in change fulfillment durations. [14].

Our work serves to filling the gap in work addressing the formal quantification of service change risk to active and depending business processes, enabling the scheduling of service changes with minimum total expected costs.

III. SERVICE TRANSITIONS AND ASSOCIATED RISK ON BUSINESS PROCESSES

The goal of service transition management is to plan and control service changes and deploying changed service releases into the production environment successfully, i.e., with minimum negative impact to the business. We assume that a service is down during the change fulfillment period. As described in Section 1, service transition in Service-Oriented Architectures is coupled with exceptionally high risk and complexity, as there are multiple interdependencies and uncertainties and many business processes might depend on a service. To estimate the risk of services changes to the business (processes), a clear picture and a formal description of the business process and service dependency structure is mandatory. Existing models like the Web Services Business Process Execution Language (WS-BPEL) or Business Process Modeling Language (BPMN) can be used to derive the dependency structure, but address many aspects not of immediate interest here. Although there is a clear mapping onto SOA models, we will now introduce a notation that is used throughout this paper to formalize the process and service dependencies relevant for our decision making.

Let I be the total number of different types of business processes i ($i=1, \dots, I$), requested stochastically following a demand distribution or profile D_i . In other words, there are I different business process definitions existing, instantiated on request. A second layer service definition j ($j=1, \dots, J$) describes an aggregated or composite service on the layer below the business process layer (i.e., the first layer). This layer represents typically automated workflows that merely string together several atomic services. Furthermore, an assignment variable u_{ij} indicates that a business process i implements service j in step u_{ij} . Steps of a business process i are enumerated by n_i ($n_i=1, \dots, N_i$). We set $u_{ij}=0$ if a business process i definition does not implement service j . In the same manner we model the dependencies of lower-level services. We enumerate the service descriptions on the next lower

aggregation level by k ($k=1, \dots, K$) and assign these third-level services by setting u_{jk} correspondingly to the step n_j ($n_j=1, \dots$

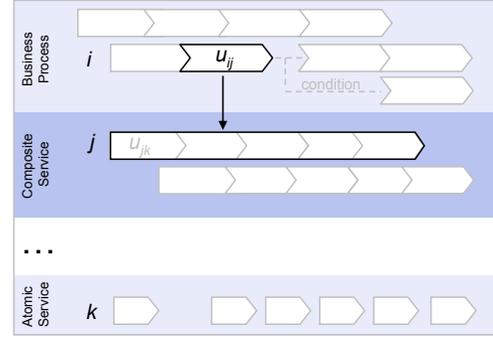


Figure 2. Three Layer Service Dependency Model

N_j) in the j service flow definition. Likewise, we set $u_{jk}=0$ if k is not implemented by j . Fig. 2 illustrates the resulting dependency structure.

Using this dependency model, one can automatically derive which higher-level services and business processes are affected by a specific service downtime.

However, to estimate the business impact of a change, additional information is required, like how many instances of business processes are affected, and how many service level agreements of these processes are expected to be violated.

The amount of affected business process instances depends on the business process demand at and before the time a change is fulfilled. Business forecasting techniques are used to estimate the demand for a certain business process during a particular period of time. With D_i as a business process i 's demand distribution profile (i.e., the demand distributions profile of all considered time slots t , D_{it} , demand forecasts d_{it} are possible for a certain time slot t (for example by setting d_{it} to D_{it} 's mean value).

For the sake of computational efficiency, we divide time into small discrete time slots, wherein we assume a fixed level demand profile. The costs of business process disruptions or delays are defined in SLAs. A SLA typically includes a process' maximum response or execution time L_i and the definition of (monetary) penalties p_i to pay on SLA violations. Depending on a SLA, penalties are paid per maximum response time violation, if the number of service level violations during a given time span exceeds a defined threshold value, or other individual agreements. Simply multiplying the number of process instances expected during the duration of a change with the penalties would overestimate change related costs, as not all running business process instances will be disrupted or delayed. For example, business process instances which already passed the step implementing the service that is going to be changed will not be affected at all, nor is there an impact on running processes instances which will execute the changed service after the change is fulfilled and the service is again available. Furthermore, business processes and services might be queued. If the time buffer, i.e., the difference between the maximum execution time and the normal or usual

execution time is large enough, there is a chance to still execute affected processes instances in a SLA compliant way.

In the following, a procedure is described to estimate the amount of SLA violations if queuing is not possible. Furthermore we extend our discussion by including queuing processes and services. We start out with a deterministic model by assuming complete knowledge of process demand per time slot and change related downtime followed by introducing a probabilistic model to account for uncertainty in both demand and service downtime.

A. SLA violations without queuing

Consider a request for change (RFC) for service j , where j will be unavailable for a duration Δt_j^{down} after the change start time t_j . The task is to estimate d_{ijt}^p , the number of SLA violations of dependent business process instances. Given this number for each affected business process, the estimated costs of changing j in t , c_{jt} are

$$c_{jt} = \sum_i p_i d_{ijt}^p \quad (1)$$

To predict d_{ijt}^p we proceed as follows: all service instances executing j during time period $[t_j; t_j + \Delta t_j^{down}]$ are disrupted. From a planning perspective, we assume equal arrival rates of business process requests (principle of indifference) as there is only aggregated knowledge of service demand per time slot available. This assumption is tight as long as the forecasting time periods are kept small. Of interest is the demand for a business process i not only during the change downtime Δt_j^{down} but also before t_j as running process instances starting before t_j might reach j during $[t_j; t_j + \Delta t_j^{down}]$. Depending on the step in which a business process i implements service j , business process instances starting after $t_j - L_i$ might be affected if j is executed in the last process step ($u_{ij} = N_i$). If j is executed in the next to last step ($u_{ij} = N_i - 1$), only process instances starting after $t_j - L_i + L_{N(i)}$ are affected, etc. On the other side, if i implements j in step N_i and the total execution duration of preceding process steps exceeds j 's downtime, instances starting during $[t_j; t_j + \Delta t_j^{down}]$ are not affected by the current change. To approximate the demand for a business processes i with j execution overlapping with $[t_j; t_j + \Delta t_j^{down}]$, d_{ijt}^p , we therefore consider business processes demand during

$$[t_j - L_i + \sum_{j'} L_{j'}; t_j + \Delta t_j^{down} - \sum_{j''} L_{j''}] \quad (2)$$

where j' is a service executed in a process i steps preceding j 's implementation step and j'' is a service executed in a step after j 's implementation step.

An alternative, more coarse-grained way of approximating d_{ijt}^p , with no further knowledge of the concrete step a process i implements a service j is described in the following: assuming an equal demand distribution around t_j , the percentage of i business process instances executing j during in $[t_j; t_j + \Delta t_j^{down}]$ is (on average)

$$\frac{L_j}{L_i} \quad (3)$$

where L_j is the execution duration of j , and L_i is the overall process execution duration. The probability that a running process instance (executing a step preceding u_{ij}) will reach j in $[t_j; t_j + \Delta t_j^{down}]$ is

$$\frac{\Delta t_j^{down}}{L_i} \quad (4)$$

Therewith, the expected total costs of SLA violations caused by changing j in t_j are

$$c_{jt} = \sum_{i: u_{ij} > 0} \left(\frac{\Delta t_j^{down} + L_j}{L_i} \right) d_{ijt} \Delta t_j^{down} p_i \quad (5)$$

B. SLA violations with queuing

We will now look at the estimated costs of changing j in time slot t if queuing (or buffering) is allowed. Here, not all business process instances executing j overlapping with $[t_j; t_j + \Delta t_j^{down}]$ are disrupted as instances can re-execute j after the change is fulfilled. If a SLA is violated depends on a process' time buffer b_i ($b_i = L_{i,max} - L_i$), where $L_{i,max}$ is the maximum execution time of a process, and L_i is the normal or usual execution time of a process. Again, the probability of a process instance currently executing j is shown in equation (3). If $b_i \leq \Delta t_j^{down}$, all considered process i instances will exceed the maximum response time. If $b_i > \Delta t_j^{down} + L_j$, no service instance is disrupted. If $\Delta t_j^{down} < b_i < \Delta t_j^{down} + L_j$, there is a change of a rollback and re-execution without SLA violation if the time buffer exceeds the amount of time already spend executing j before t_j plus j 's downtime Δt_j^{down} . This probability is shown in equation (6)

$$\left(\frac{L_j}{L_i} \right) \left(1 - \frac{b_i}{L_i} \right) \quad (6)$$

The probability that a running process instance (executing preceding steps) will reach j in $[t_j; t_j + \Delta t_j^{down}]$ is shown in (4). If $b_i > \Delta t_j^{down}$, all services are delivered successfully. If $b_i < \Delta t_j^{down}$, the average rate of successful delivered business process instances is

$$\left(\frac{\Delta t_j^{down}}{L_i} \right) \left(\frac{b_i}{\Delta t_j^{down}} \right) \quad (7)$$

C. Non-Linear Business Processes and Service Flows

The estimation of change related penalties as introduced in the previous section assumes linear business processes and service flows with a predetermined sequence of service executions. In practice, business processes might take different

branches or service flow paths based on certain conditions. One branch might include a service to be changed while others do not. Hence, business process forecasting ignoring such conditional branches overestimates the number of SLA violations and costs. A finer-grained demand forecast is required for each possible branch. This forecast can be derived by analyzing the history of the different executed branches in the same way the total demand for linear processes is derived by business forecasting methods. We model each branch as own business process as shown in Fig. 3.

Using this statistical means, one can model forked business processes. Processes including iterative sequences like loops can be demodulated in the same manner, by defining each possible flow as an own process and by assigning probabilities derived from statistical analyses of log data.

IV. CHANGE SCHEDULING DECISION MODELS

We will first introduce a basic change scheduling decision model for shared services underlying a number of restrictive assumptions like perfect knowledge of business process demand per time slot and deterministic change related downtimes of services. Afterwards, we will propose model variants considering uncertainty in business process demand and stochastic service downtime. Based on these model formulations, a couple of extensions are introduced to consider other types of operational risks and costs associated with service transitions. Furthermore, we address the problem of handling correlated changes.

A. Basic Deterministic Model

We will now introduce a deterministic mathematical programming model (DMP) to solve the problem of finding the schedule for a set of uncorrelated changes J_{RFC} with minimum overall service level violation costs in environment without queuing. Business process demand per time slot t , d_{it} , the downtime of a service after the change start time, Δt_j^{down} , and execution durations of services, L_j , are approximated by using their mean values. A penalty is paid per SLA violation.

We introduce a binary decision variable $x_{j,t} \in \{0,1\}$ indicating whether j 's change is started in t_j or not.

The objective functions to minimize the total sum of penalties resulting from changes in service infrastructures without queuing is

$$\min \sum_{j \in J_{RFC}} \sum_{i: u_{ij} > 0} \sum_t \left(\left(\frac{\Delta t_j^{down} + L_j}{L_i} \right) d_{ijt} \Delta t_j^{down} \right) p_i x_{j,t} \quad (8)$$

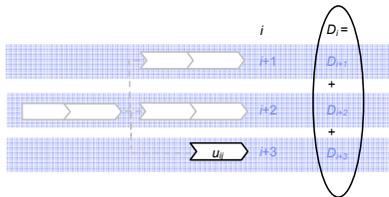


Figure 3. Non-Linear Business Processes and Service Flows

We set the beginning of our change planning period to $t=0$ and assume to obtain J_{RFC} before $t=0$ (Note that in practice, changes will be requested on a continuous time base rather than bundled. The usual way to proceed is to re-calculate the optimization problem each time a new RFC is submitted. More advanced methods might forecast aggregated RFC 'demand' if changes are submitted in regular sequences. As we divide time into discrete time slots, time related parameters are of positive integer type ($t_j, \Delta t_j^{down}, b_i, L_i, L_j \in \mathbb{Z}_0^+$) and penalties and demand parameter are of positive real type ($d_{it}, p_i \in \mathbb{R}_0^+$).

As further constraints we introduce change related deadlines t_j^d . Depending on the severity of a change, there is generally a priority associated with a change, defining a deadline when a change needs to be implemented. This constraint can be formulated as

$$\sum_{t_j + \Delta t_j^{down} < t_j^d} x_{j,t} = 1, \quad \forall j \in J_{RFC} \quad (9)$$

Note that a change deadline is originally defined as a period Δt_j^d after t_j^{RFC} , the time the RFC for j arrives. As we define $t_j^{RFC} = 0$, setting the deadline to t_j^d instead of $t_j^{RFC} + \Delta t_j^d$ suffices in our case.

B. Stochastic Change Scheduling Model

So far, we used deterministic approximations for expected demand, service downtime and service execution durations.

One should expect that ignoring the probabilistic nature of demand, downtime and execution time has a negative impact on the decision making. Suppose a service j change, and a depending business process i with extremely high penalties to pay on service level violations. The average change related downtime of j is 10 but varies broadly, and the decision is either to start the change in $t=0$ or in $t=50$. The demand for i is expected to be slightly lower during $t=0-9$ than during $t=50-59$ but increases rapidly from $t=10$ on, while demand is expected to be of constant level after $t=59$. The deterministic model would certainly select $t=0$ while a stochastic model explicitly taking into account uncertainty of downtime would select $t=50$, which would be the better decision.

However, putting too much stochastic information into a decision model makes it – at least for medium and large problem sizes – intractable due to the large number of resulting decision variables and limits therefore its practical applicability.

Therefore, we draw on a stochastic programming formulation with simple recourse as introduced for example by Birge and Louveaux to consider the stochastic nature of the variables while keeping the model computable [15, 16].

This is illustrated using a change related downtime probability distribution as shown in Fig. 4. We separate the distribution into N sequential discrete sections n ($n=1, \dots, N$). The cumulated probability (integral) of a section is then interpreted as the downtime probability of one dedicated time slot in the section, while we suppose the downtime can only take these discrete downtime values: $\Delta t_j^{down} \in \{\Delta t_{j,1}^{down}, \Delta t_{j,2}^{down}, \dots, \Delta t_{j,N}^{down}\}$.

$\dots, \Delta t_{j,N}^{down}\}$. The resulting objective function can be formulated as

$$\min \sum_{j \in J_{RFC}} \sum_{t: u_{jt} > 0} \sum_{i=1}^N P(\Delta t_{j,n}^{down}) \left(\frac{\Delta t_{j,n}^{down} + L_j}{L_i} \right) d_{ijt} \Delta t_{j,n}^{down} p_i x_{j,t} \quad (10)$$

The right part of the objective function computes the costs that would be resulted if the downtime would have been exactly $\Delta t_{j,n}^{down}$; the term on the left is a correction for the uncertainty in downtime (a weight).

Likewise, we model the other stochastic variables like business process demand during a time slot or the execution time of a service. Note that usually the parameters or even the type of distributions will depend on which time slot you consider.

C. Change Fulfillment Deadlines and Waiting Costs

As already mentioned, a change needs to be fulfilled in a maximum change fulfillment time Δt_j^d after a change request is submitted. As discussed previously in this paper, the urgency depends on the priority of a change. In the basic deterministic model formulation we assumed that this deadline is mandatory.

Considering the uncertainty in the time needed to perform the service change (we assume the service to be down during change activities) it can no longer be guaranteed to fulfill a change before the agreed change deadline; only a probability can be assigned to fulfilling the change in time. Therefore, the restriction that a change needs to be fulfilled before t_j^d of the change deadline needs to be relaxed to

$$\sum_t x_{j,t} = 1, \forall j \in J_{RFC} \quad (11)$$

Exceeding a change deadline might entail a predefined penalty and extra payments for each additional time slot needed to fulfill the change. The later a change is started, the higher the expected costs of a deadline violation will be, since the probability of completing change implementations before their deadline will decrease continuously.

Let the fixed penalty on change deadline violation be α , and the additional costs per time slot a deadline is exceeding be β .

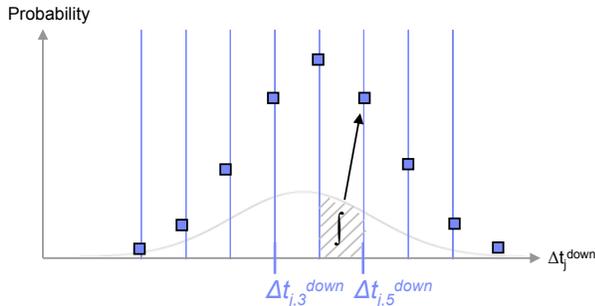


Figure 4. Probabilistic Modelling, Downtime

Therewith, the expected overall deadline violation cost function which needs to be added to the objective function as formulated in our decision model is

$$\min \sum_t (\alpha(t_j + \Delta t_j^{down} - t_j^d) > 0) + \beta(\max(0, t_j + \Delta t_j^{down} - t_j^d)) x_{jt} \quad (12)$$

Note that for reasons of brevity we provide equations with only the service downtime modeled stochastically while other stochastic parameters are approximated by their mean values.

Furthermore, the moment an RFC is submitted, there may already be a need felt for the change to be implemented as the business may suffer until the change has been fulfilled; for example, this may be due to a service being unavailable as would happen if the change request was initiated as a result of an incident, or there may be other negative impact causes, like lost opportunities such as would occur for a change meant to bring up a new required service. With γ as the implicit costs of waiting one more timeslots for a change to be fulfilled, the waiting costs can be formulated as

$$\sum_t \gamma(t_j + \Delta t_j^{down}) x_{j,t} \quad (13)$$

D. Allowed Change Windows

Furthermore, the fulfillment time of a change might be restricted to a number of allowed change window time slots, e.g. at weekends or during night times. Violating a change window restriction might have serious impact on the business, as that would mean a service is down in times this service is frequently required. Therefore, penalties might result from exceeding a change window l ($l=1, \dots, L$). Let T_{cj} ($T_{cj} = \{t_{cj}^{start}, \dots, t_{cj}^{end}\}, \dots, \{t_{cjl}^{start}, \dots, t_{cjl}^{end}\}$) be the set of allowed change windows. As change related downtime might be of uncertain length, there is an increasing risk of violating the change window constraints the later a change is started. With δ as the costs per time slot a change window is exceeded, and the restriction that a change has to (at least) start inside a change window ($t_j \in T_{cj}$), the part that has to be added to the objective function as formulated in our decision model is

$$\min \sum_t \max(0, \delta(t_j + \Delta t_j^{down} - \min(t_{jl}^{end}, t_{jl}^{end} > t_j))) x_{j,t} \quad (14)$$

E. Correlated Changes

The basic model formulation handles multiple independent changes. To schedule changes in a mandatory order, a constraint for each dependency has to be added to the decision model formulation. Firstly, changes might need to be started in a certain sequence ($t_j < t_{j+1} < t_{j+2} < \dots$) or a change must be fulfilled before the next change may get scheduled ($t_j + \Delta t_j^{down} < t_{j+1} + \Delta t_{j+1}^{down} < \dots$). The constraints in our mathematical model formulation are therefore $x_{jt} < x_{(j+1)t} < t_{(j+2)t}$, or $x_{jt} + \Delta t_j^{down} < x_{(j+1)t} + \Delta t_{j+1}^{down} < t_{j+2}$, respectively.

Besides mandatory change scheduling orders, changes might be correlated for example in terms of a reduction of aggregated downtime when executing changes together (imagine two changes to a server operating system, both requiring a reboot. The overall change duration might be reduced by applying these changes together, but this may result in higher risk in terms of higher downtime variance (incompatibilities, etc.).

While arbitrary statistical values can be chosen, in our example we focus on mean (M) and variance (V) deviation. Therefore, we consider two changes to j and $j+1$ as correlated if either

$$M(\Delta t_j^{\text{down}}(t) + \Delta t_{j+1}^{\text{down}}(t)) \neq M(\Delta t_j^{\text{down}}(t) + \Delta t_{j+1}^{\text{down}}(t + \Delta t)) \text{ and/or}$$

$$V(\Delta t_j^{\text{down}}(t) + \Delta t_{j+1}^{\text{down}}(t)) \neq V(\Delta t_j^{\text{down}}(t) + \Delta t_{j+1}^{\text{down}}(t + \Delta t))$$

We treat each change item combination with significant deviant aggregated statistical mean and/or variance values as one single change. The decision to make is to either schedule all included single changes separately or to schedule the novel ‘aggregated’ change instead. This XOR constraint can be formulated as follows (if the question is to either change j and $j+1$ separately, or, alternatively the aggregated change ($j, j+1$))

$$\sum_t x_{j,t} + x_{(j+1),t} + 2x_{(j,j+1),t} = 2 \quad (15)$$

Furthermore, the change deadline for ($j, j+1$) is set to $\min(t_{j,RFC} + \Delta t_j^d, t_{j+1,RFC} + \Delta t_{j+1}^d)$.

F. Change Re-Scheduling

The decision model selects the time slot with the lowest expected overall costs based on business process demand forecasting. However, when approaching to t_j , further knowledge is available of process demand and process instances’ states (progress). This knowledge can be used to reschedule the change start time t_j . For example, if in $(t_j - 1)$ more business process instances are running than expected, or a higher percentage of running instances is currently executing service j , there is a decision to make on whether to retain t_j or to wait several timeslots. However, increasing delay costs and a higher probability of violating change window restrictions have to be taken into account when making such a decision. Note that demand forecasting for processes might be adapted by using short term prognoses if current demand differs significantly from demand expected beforehand. Furthermore, business process request arrivals might be modeled as Poisson Process to consider the uncertainty regarding the exact arrival rates, with $P_{\lambda(i)}(r=k)$ as the probability of k incoming service i requests in t . As we did with downtime uncertainty, we model the impact of different possible arrival rates weighted by their probabilities.

V. EXPERIMENTAL ANALYSIS

In this section, we analyze and discuss the efficiency of the scheduling models proposed in this paper. In our experimental evaluations we compared variants of our models to the optimal solutions (by scanning the total solution space), with total

change related costs under different service infrastructures, demand scenarios, and downtime distributions used as a benchmark. First, the experimental set-up that we used for our preliminary experiments is described, and second, we report the results of our experiments and discuss their outcome.

A. Experimental Set-Up

We analyzed 12 different service infrastructure scenarios under different business process demand profiles. We used real-world data and data based on patterns found in literature to generate these infrastructure scenarios and demand profiles.

The durations of each experiment was set to 300 time slots t ($t = 0, \dots, 299$), where the length of each time slot was set to one hour. The change deadline was set to $t_j^d = 275$ with fixed costs if this restriction was violated and additional costs per exceeded times slot. In our first evaluations, change windows, and waiting costs were not considered. To allow for sensitivity analysis how variations in the output of our models can be apportioned to variations of j ’s downtime distribution, we repeated each experiment until our results were significant (referred to as experimental item, average over all outcomes) for each downtime distribution. We analyzed 8 different downtime distributions with increasing variance. To configure and automate our experiments and to analyze our experimental outcomes a simulation tool has been developed (see Fig. 5). The figure shows a visualization of an example service infrastructure scenario used in our experiments with two business processes, a linear and a forked process.

An example business process demand scenario is shown in Fig. 6. The graph shows the mean demand level M per time slot. We adapted the demand level after each time slot to generate a demand profile following these curves. During a time slot, we generated demand following a $(M, 0.20M)$ normal probability distribution (uniformly distributed).

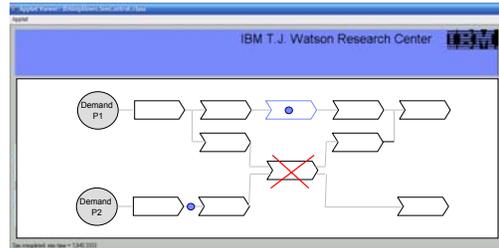


Figure 5. Analyzed Service Infrastructure Scenario

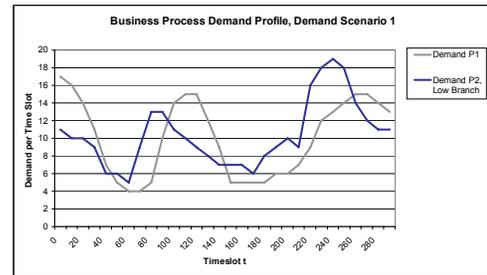


Figure 6. Example Business Process Demand Scenario

B. Experimental Results

Experimental results show that the probabilistic decision model with a simple resource of the service downtime distributions (applying the objective function as shown in equation (10)) found the optimal solution for all experimental items. In experiments with low service downtime variance (less than 15% of the mean downtime duration), the deterministic model selected the change start time slot with minimum costs. Except one demand scenario with almost flat process demand levels, the deterministic variant never found the optimal solution in scenarios with one of the two highest downtime variances. Fig. 7 presents aggregated results of the cost savings by using either the deterministic or the probabilistic scheduling model. The bars show the change related costs when using one of the two decision model variants relative to the average costs over all scenarios (with a certain downtime variance level) when the change start time was selected randomly.

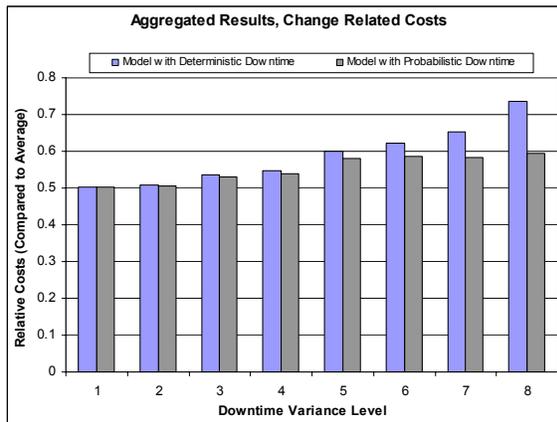


Figure 7. Aggregated Experimental Results

VI. SUMMARY AND OUTLOOK

In this paper we introduced a model to analyze the business impact of changes in a network of services. We analyzed change related operational risks on active business process instances and techniques to relate these risks to financial metrics.

To our best knowledge, our work is the first to formally quantify the risk of changing services in SOA environments to the business (processes), or that derives decision models which allow organizations to schedule service changes with minimum total expected costs.

In our experimental analyses we evaluated the efficiency of our models compared to the optimal and average solution, with total change related costs under different demand scenarios and downtime distributions used as a benchmark. We conducted numerical experiments with various business process demand scenarios and different downtime distributions and made initial efficiency statements. Experimental results show that the proposed probabilistic model derived the optimal solution in all of our experiments, and the deterministic model only if the downtime variance was low. We hoped to obtain such results, but, however, it is not obvious that taking more stochastic

information into account automatically leads to better results (e.g., in Airline Revenue Management, most of the deterministic seat inventory control approaches perform better than their stochastic pendants).

Future working plans are more exhaustive sets of experiments with different, possibly real-world, business scenarios. Furthermore, we work on meta-heuristics and pre-selecting time-slots to solve much large problem sizes. We intend to additionally explore the impact of rescheduling change times when approaching the planned change start time, the impact of uncertain service execution durations, and the impact of latency and change window violation costs. We also plan to test the models in the field as a decision support tool for change scheduling in selected businesses. Finally, we plan to tighten the connections to web services standards for describing the SOA network of services including IT resources.

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